Lab 6: Unconditional Quantile Regression Econometrics Beyond Ordinary Least Squares

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Unconditional Quantiles: Quantiles of the overall distribution.



Figure 1: Conditional and Unconditional Quantiles

Conditional and Unconditional Quantile Regression Coefficients

CQR coefficients: Workers' benefit of **transferring into** public sector jobs. UQR coefficients: Earnings increase if **percentage of public sector workers** increases.



Figure 2: Fournier, 2012, Fig 3: "Conditional and unconditional quantile regression estimates of the impact on earnings from working in the public sector"

Changes in unconditional quantiles \Rightarrow retrieve changes in quantile ratios (eg. 80/20, 90/10).

Extend UQR to other distributional properties \Rightarrow retrieve changes in Gini coefficient, mean-median ratio.

In OLS $\hat{\beta} = E(Y \mid X = 1) - E(Y \mid X = 0)$, the conditional effect of transferring between groups, as well as $\hat{\beta} = \frac{\partial \mu(p)}{\partial p}$ with $p = \frac{\sum_{i=1}^{N} 1(X=1)}{\sum_{i=1}^{N} 1(X=1) + 1(X=0)}$.

This does not hold for CQR, usually $\hat{\beta}_{\tau} = F_Y^{-1}(\tau \mid X = 1) - F_Y^{-1}(\tau \mid X = 0) \neq \frac{\partial q_{\tau}(p)}{\partial p} = P(Y > q_{\tau} \mid X = 1) - P(Y > q_{\tau} \mid X = 0).$

"The influence function $IF(Y; v, F_Y)$ of a distributional statistic $v(F_Y)$ represents the influence of an individual observation on that distributional statistic." (Firpo et.al., 2009, 954)

Re-centered influence function RIF: $RIF(Y; v, F_Y) = v(F_Y) + IF(Y; v, F_Y)$. Note: $E[RIF(y; v, F_Y)] = E[v(F_Y)]$.

$$IF(Y;q_{\tau},F_Y) = \frac{\tau - 1(Y \le q_{\tau})}{f_Y(q_{\tau})}$$
(1)

$$RIF(Y; q_{\tau}, F_Y) = q_{\tau} + IF(Y; q_{\tau}, F_Y)$$
(2)

RIF regression model: conditional expectation of the RIF $E[RIF(Y; q_{\tau}, F_Y) | X]]$. RIF regression: OLS estimation $RIF(Y; q_{\tau}, F_Y) | X_i] = \beta_{RIF}X_i + \epsilon_i$. $\hat{\beta}_{RIF}$ corresponds to the **effect of a marginal change in** X on the unconditional

 β_{RIF} corresponds to the effect of a marginal change in X on the unconditional quantile of Y.

Necessary steps: Estimate quantiles $q_{\tau}(Y)$, density $f_Y(q_{\tau})$ (eg. by Kernel estimation), calculate dummy variable $1(Y \leq q_{\tau})$ (trivial).

For a stochastic experiment with outcome variable Y, the **probability density** function (PDF) gives the probability of a certain realization Y = y to be observed.



Figure 3: Probability Density Functions of a Discrete and a Continuous Variable Y $\mathsf{N}(0,1)$

Distributions: Cumulative Density Function F(Y)

The cumulative density function (CDF) gives the probability that a random variable **realizes below a treshold level** P(Y < y). It is 0 for the minimum and 1 for the maximum range of Y.

For observations it can be understood as the fraction of the population with realizations below some observation.



Figure 4: Cumulative Density Functions of a Discrete and a Continuous Variable Y N(0,1)

Joint Distributions

The **joint distribution** $f_{X,Y}(x, y)$ of two variables X, Y gives the probability of observing two values X = x, Y = y at the same time.



Figure 5: Wikipedia: "Many sample observations (black) are shown from a joint probability distribution. The marginal densities are shown as well."

A functional $v(F_Y)$ projects from a function F_Y to the space of real numbers \mathbb{R} , ie. $v : F_v \to \mathbb{R}$, eg. the mean.

Hampel, 1968: The influence function gives the infinitesimal behavior of a functional v.

One can have two distributions of the same class F_Y and G_Y (eg. one N(0,1) and one N(1,1)).

Then there exists a **mixing distribution that is** t **units away** from F_Y in the direction of G_Y : $F_{Y,t:G_Y} = (1-t)F_Y + tG_Y = t(G_Y - F_Y) + F_Y$.

$$\lim_{t\downarrow 0} \frac{v(F_{Y,t\cdot G_Y}) - v(F_Y)}{t} = \frac{\partial v(F_{Y,t\cdot G_Y})}{\partial t}|_{t=0}$$
$$= \int IF(y;v,F_Y) \cdot d(G_Y - F_Y)(y)$$
(3)

The von-Misès linear approximation of $v(F_{Y,t}, G_Y) - v(F_Y)$:

$$v(F_{Y,t\cdot G_Y}) - v(F_Y) = v(F_Y) + \int IF(y;v,F_Y) \cdot f(G_Y - F_Y)(y) + r(t;v;G_Y,F_Y)$$
(4)

Neutralize the "remainder term" r(.) (by setting $G_Y = \Delta_y$ and t = 1), for the **re-centered influence function RIF**

$$RIF(y; v, F_Y) = v(F_Y) + \int IF(s; v, F_Y) d\Delta_y(s) = v(F_Y) + IF(y; v, F_Y)$$
 (5)

Key result: "the impact of a marginal change in the distribution of X on $v(F_Y)$ can be obtained using the conditional expectation of the $RIF(Y; v, F_Y)$ ". (Firpo et.al., 2009, 957)

$$RIF(Y; q_{\tau}, F_{Y}) = q_{\tau} + IF(Y; q_{\tau}, F_{Y})$$
$$= q_{\tau} + \frac{\tau - 1(Y \le q_{\tau})}{f_{Y}(q_{\tau})}$$
(6)

Bivariate RIF



$$RIF(y_i; q_\tau, F_Y) \sim \alpha + X'_i \beta + \epsilon_i$$

$$\beta = \frac{\partial RIF(y_i; q_\tau, F_Y)}{\partial X_i}$$
(8)

```
df_tmp %>%
    lm(rif01 ~ x, data=.)
```

```
Call: Im(formula = rif01 \sim x, data = .)
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Coefficients: (Intercept) × -31.718 2.902

- Same regression between two groups: different endowments and coefficients (eg. women have more years of education, but smaller income returns for each year).
- Oaxaca (1973) and Blinder (1973): Decompose effects by calculating counterfactual distribution if group A had endowments of group B.

$$\begin{aligned} \Delta \bar{Y} &= \bar{Y}_A - \bar{Y}_B \\ \Delta \bar{Y} &= \bar{X}'_A \hat{\beta}_A - \bar{X}'_B \hat{\beta}_B \\ \Delta \bar{Y} &= (\bar{X}_A - \bar{X}_B)' \hat{\beta}_B + \bar{X}'_B (\hat{\beta}_A - \hat{\beta}_B) + (\bar{X}_A - \bar{X}_B)' (\hat{\beta}_A - \hat{\beta}_B) \end{aligned} \tag{9}$$

- UQR estimated by OLS: Oaxaca-Blinder decomposition is trivial.
- Interpretation of Coefficients is not.
- In surveys, re-weighting becomes crucial.