10: Co-Integration in Panels GECO 6281 Advanced Econometrics 1 (Lab)

Patrick Mokre

Fall 2020

- Panel Data: LSDV, Fixed Effects and Random Effects
- Dynamic Panel Data & Instrumentalization: Anderson-Hsiao, Arellano-Bond, Arellano-Bover
- Auto-Regressive and Moving Average Processes
- Stationarity in Time Series and Panels
- Vector Error Correction (VECM) and Auto-Regressive Distirbuted Lag (ARDL)

Cointegration

Two time series x_t and y_t are cointegrated if they are both I(1), but a linear combination $\lambda_1 x_t + \lambda_2 y_t$ of the processes is I(0) stationary.

Simple Test: Regress y_t on x_t , estimate residuals \hat{e}_t . Then perform an augmented Dickey-Fuller test on \hat{e}_t by regressing its first difference on the lagged level. The **ADF** test statistic then corresponds to the t-statistic on the corresponding coefficient.

$$y_t = \alpha + x_t'\beta + \epsilon_t \tag{1}$$

$$\Delta \hat{\epsilon}_t = a + \rho \hat{\epsilon}_{t-1} + u_t \tag{2}$$

$$ADF\epsilon_t = \frac{\hat{\rho}}{s.e.(\hat{\rho})} \tag{3}$$

Cointegration is a test on non-stationary time series to establish (or reject) a "long-run" relationship between the two.

First difference time series tend to "wander". Co-Integration implies that series are wandering together in long-run equilibrium (although groups of them can wander arbitrarily).

- Pure time series application, appropriate for vector error correction (VEC).
- i.e. not a panel cointegration test, but links time series and panel cointegration testing.
- Solves the problem of how many co-integrating relationships to include in the VEC.
- Let r be the number of co-integrating, n the number of all processes. H0: r < n, $H_A: r = n$.
- Sequential testing for r = (1, 2, N), the first non-rejection indicates the number of co-integrating processes.
- Method: Likelihood-Ratio test.

Logic (one variable vector):

$$\begin{split} y_t &= \phi y_{t-1} + e_t \mid -y_{t-1} \\ \Delta y_t &= (\phi-1)y_{t-1} + e_t \end{split}$$

For the unit root case $\phi = 1 \Rightarrow (\phi - 1) = 0$.

Johansen Unit Root Test for Cointegration 2

Let $\boldsymbol{y}_t = (y_{1,t}, y_{2,t}, ..., y_{N,t})$ be a vector.

$$\begin{split} y_t &= \Phi y_{t-1} + e_t \\ \Delta y_t &= \Phi y_{t-1} - y_{t-1} + e_t \\ &= (\Phi - I) y_{t-1} + e_t \end{split}$$

The rank of $(\Phi - I)$ is equal to the number of co-integrating, non unit-root, processes: Johansen Rank Test. The rank of a matrix is the maximum of linearily independent column or row of a matrix (definitions are equivalent).

Both a constant term $A_0 = (a_{1,0}, a_{2,1}, ..., a_{N,0})$ and a drift term Σ can be included. Let D_t be a matrix of yearly dummy variables. Let p be some appropriate lag length.

Formulation as long-run vector error correction model:

$$\Delta Y_t = A_0 + \Sigma D_t + \Pi A_{t-p} + \sum_{i=1}^{p-1} \Gamma_i \Delta Y_{t-i} + e_t \text{ where } \Gamma_i = \sum_{j=1}^i [\Pi_j] - I, i = 1, ..., p-1$$

Formulation as a transitory VECM:

$$\Delta Y_t = A_0 + \Sigma D_t + \sum_{i=1}^{p-1} \Gamma_i \Delta Y_{t-i} + \Pi Y_{t-1} \text{ where } \Gamma_i = (\Pi_{i+1} + \ldots + \Pi_p), i = 1, \ldots, p-1$$

The long-run and the transitory VECM are equivalent. In both cases, testing revolves around $\Pi=\Pi_1+...+\Pi_p-I.$

The test finds the correct rank of Π , where $rank(\Pi)=0$ indicates a regular first difference VAR relationship, and $rank(\Pi)=n$ indicates a VECM where all vectors are mutually cointegrated.

The STATA command for Johansen testing is vecrank.

Co-Integration in panels refers to a **long-run relationship** between different variables that holds **over cross-sectional units**.

Co-Integration in Panels is to be estimated when **non-stationarity** cannot be rejected. For stationary series, there is better estimation techniques (VAR, ECM, VECM, ...).

The problem with early co-integration testing was the **weak power of the tests**, i.e. a small likelihood to reject the null hypothesis. This seems to be due to the (short) length of the series. Implementing panels rather than extending the length of the series has numerous advantages. (Pedroni 2004)

$$y_{i,t} = \alpha_i + \delta_i t + \beta_i X_{i,t} + e_{i,t} \tag{4}$$

- \blacktriangleright $y_{i,t}$ and $X_{i,t}$ are I(1) for each panel member i.
- Fixed effects α_i , deterministic trends δ_i and co-integration vectors β_i vary between members.
- H₀: None of the panel members are co-integrated.
- $H_{A,1}$ All members are co-integrated with the same co-integration vector. (Kao 1999)
- $H_{A,2}$: All members are cointegrated (the co-integration feature has to be the same for all β_i). (Pedroni 2004)
- \blacktriangleright $H_{A,3}$: A significant portion of the individuals are co-integrated. (Pedroni 2004)

For homogenous slope parameters β and strictly exogenous regressors, no co-integration is asymptotically equivalent to raw panel unit root tests.

For homogenous slopes and partially endogenous regressors, critical values for the unit root tests have to be adjusted.

For heterogenous slope parameters, residual-based test for co-integration must not be based in **pooling** test statistics.

$$(y_{i,t}, X'_{i,t}) = z'_{i,t} = z'_{i,t-1} + \xi_{i,t}$$
(5)

$$\xi_{i,t}' \equiv (\xi_{i,t}^y, \xi_{i,t}^x) \tag{6}$$

Assumptions: Let $B_i\Omega_i$ be a Brownian motion with asymptotic variance Ω_i .

A Brownian motion is a process with small, random movements in any dimension.

 ¹/_{√T} ∑_{t=1}[Tr] → B_iΩ_i Central Limits Theorem hold for each individual member.

 E[ξ_{i,t}ξ'_{js}] = 0 ∀i ≠ j, s ≠ t Central Limit Theorem holds in the cross-sectional dimension. \end{align*}

Co-Integration with homogenous slopes (Kao)

$$\begin{aligned} x_{it} &= x_{i,t-1} + e_{i,t} \\ y_{i,t} &= y_{i,t-1} + v_{i,t} \\ y_{i,t} &= \alpha + \beta x_{i,t} + u_{i,t} \\ \hat{u}_{i,t} &= \rho \hat{u}_{i,t-1} + \epsilon_{i,t} \end{aligned} \tag{7}$$

Dickey-Fuller Test:

$$\hat{\rho} = \frac{\sum_{i=1}^{N} \sum_{t=2}^{T} \hat{u}_{i,t} \hat{u}'_{i,t}}{\sum_{i=1}^{N} \sum_{t=2}^{T} \hat{u}^{2}_{i,t}}$$
$$DF = \sqrt{N}T(\hat{\rho} - 1)$$

Augmented Dickey Fuller Test:

$$\begin{split} \hat{u}_{i,t} &= \rho \hat{u}_{i,t-1} + \sum_{j=1}^p \phi_j \Delta \hat{u}_{i,t-j} + e_{i,t,p} \\ ADF &= \frac{\hat{\rho} - 1}{s.e.(\hat{\rho})} \end{split}$$

where the number of lags p is chosen such that the errors $e_{i,t,p}$ are serially uncorrelated.

$$\begin{aligned} x_{it} &= x_{i,t-1} + e_{i,t} \\ y_{i,t} &= y_{i,t-1} + v_{i,t} \\ y_{i,t} &= \alpha + \beta x_{i,t} + u_{i,t} \end{aligned}$$

Let Ω_i be the asymptotic covariance matrix for each member, Σ_i contemporaneous covariance among the components of $\xi_{i,t}$ for a given cross section i, and Γ_i the dynamic covariance between components of $\xi_{i,t}$. Ω gives the long run co-integrating relationship (if there is one).

$$\Omega_i = \lim_{T \to \infty} E\left[\frac{1}{T} \sum_t^T \xi_{i,t} \sum_t^T \xi'_{i,t}\right]$$
(10)

$$\Omega_i = \Sigma_i + \Gamma_i + \Gamma_i' \tag{11}$$

$$\Sigma_i = \lim_{T \to \infty} \frac{1}{T} \sum_{t}^{T} E[\xi_{i,t} \xi'_{i,t}]$$
(12)

$$\Gamma_{i} = \lim_{t \to \infty} \frac{1}{T} \sum_{k=1}^{T-1} \sum_{t=k+1}^{T} E[\xi_{i,t}\xi'_{i,t}]$$
(13)

Fortunately, $\hat{\Omega}_i$ can be efficiently estimated after de-composition. Retrieve $\hat{xi}_{i,t} = (\hat{v}_{i,t}, \hat{\epsilon}_{i,t})$ from auto-regressions for each member i:

$$\begin{split} x_{i,t} &= \rho_i x_{i,t-1} + \epsilon_{i,t} \\ y_{i,t} &= \rho_i y_{i,t-1} + v_{i,t} \end{split}$$

$$\begin{split} \hat{\Omega}_{i} &= \hat{\Sigma}_{i} + \hat{\Gamma}_{t} + \hat{\Gamma}'_{i} \\ \hat{\Omega}_{i} &= \frac{1}{T} [\sum_{t}^{T} \hat{x} \hat{i}_{i,t} \hat{x} \hat{i}_{i,t} + \sum_{s}^{k_{i}} (1 - \frac{s}{k_{i} + 1}) \sum_{t=s+1}^{T} (\hat{\xi}_{i,t-s} \hat{\xi}'_{i,t} + \hat{\xi}_{i,t} \hat{\xi}'_{i,t-s})] \end{split}$$
(14)

For the estimation, one runs the co-integration regression for each panel member, and calculates the lower triangular decomposition $\hat{L}_{11,i} = (\hat{\Omega}_{11,i} - \hat{\Omega}_{21,i}^2 / \hat{\Omega}_{22,i})$.

Then one runs the residual regression for each panel member and calculates the group mean statistics (H_0 : no cointegration).

$$y_{i,t} = \alpha + \beta x_{i,t} + u_{i,t}$$

$$\hat{u}_{i,t} = \rho \hat{u}_{i,t-1} + e_{i,t}$$
 (15)

Furthermore, one calculates the auxiliary statistic $\lambda_i = (\hat{\sigma}_i^2 - \hat{s}_i^2)$.

$$\begin{split} \hat{\sigma}_{i}^{2} &= \frac{1}{T} [\sum_{t}^{T} \hat{e}_{i,t}^{2} + 2\sum_{s}^{k_{i}} (1 - \frac{s}{k_{i} + 1}) \sum_{t=s+1}^{T} \hat{e}_{i,t} \hat{e}_{i,t-s}] \\ \hat{s}_{i,t}^{2} &= \frac{1}{T} \sum_{t}^{T} \hat{e}_{i,t}^{2} \end{split}$$

One can then calculate the group mean Z-statistics.

$$\begin{split} \tilde{Z}_{\hat{\rho},N,T} &= \sum_{i}^{N} \frac{\sum_{t}^{T} (\hat{u}_{i,t-1} \Delta \hat{u}_{i,t} - \hat{\lambda}_{i})}{\sum_{t}^{T} \hat{u}_{i,t-1}^{2}} \\ \tilde{Z}_{t,N,T} &= \sum_{i}^{N} \frac{\sum_{t}^{T} (\hat{u}_{i,t-1} \Delta \hat{u}_{i,t} - \hat{\lambda}_{i})}{\sum_{t}^{T} \frac{1}{\hat{L}_{11,i}^{2}} \hat{u}_{i,t-1}^{2}} \end{split}$$
(16)